

Linear Aspects in Random Coefficient Regression Models: Optimal Designs for Different Design Criteria

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RCR model

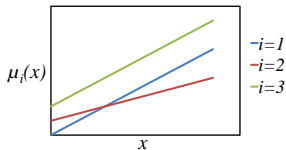
$$Y_{ij} = \mathbf{f}(x_j)^\top \beta_i + \varepsilon_{ij}, \quad i = 1, \dots, n, \quad j = 1, \dots, m, \quad x_j \in \mathcal{X}$$

- $\mathbf{f} = (f_1, \dots, f_p)^\top$
- $\beta_i = (\beta_{i1}, \dots, \beta_{ip})^\top$, $E(\beta_i) = \beta$ *unknown*, $\text{Cov}(\beta_i) = \sigma^2 \mathbf{D}$
Here: D regular
- $E(\varepsilon_{ij}) = 0$, $\text{Var}(\varepsilon_{ij}) = \sigma^2$
- *All β_i and all ε_{ij} uncorrelated*

Gladitz & Pilz (1982): β known

Individual response

$$\mu_i(x) = \mathbf{f}(x)^\top \beta_i$$



RCR model

$$Y_{ij} = \mathbf{f}(x_j)^\top \boldsymbol{\beta}_i + \varepsilon_{ij} \quad \equiv \quad Y_{ij} = \mathbf{f}(x_j)^\top \boldsymbol{\beta} + \mathbf{f}(x_j)^\top \mathbf{b}_i + \varepsilon_{ij}$$

for $\mathbf{b}_i = \boldsymbol{\beta}_i - \boldsymbol{\beta}$
 with $E(\mathbf{b}_i) = 0$ & $\text{Cov}(\mathbf{b}_i) = \sigma^2 \mathbf{D}$

i -th individual $\mathbf{Y}_i = (\mathbf{Y}_{i1}, \dots, \mathbf{Y}_{im})^\top$:

$$\mathbf{Y}_i = \mathbf{F}\boldsymbol{\beta} + \mathbf{F}\mathbf{b}_i + \boldsymbol{\varepsilon}_i$$

$$\mathbf{F} = (\mathbf{f}(x_1), \dots, \mathbf{f}(x_m))^\top \quad \& \quad \boldsymbol{\varepsilon}_i = (\varepsilon_{i1}, \dots, \varepsilon_{im})^\top$$

All individuals $\mathbf{Y} = (\mathbf{Y}_1^\top, \dots, \mathbf{Y}_n^\top)^\top$:

$$\mathbf{Y} = (\mathbf{1}_n \otimes \mathbf{F})\boldsymbol{\beta} + (\mathbf{I}_n \otimes \mathbf{F})\mathbf{b} + \boldsymbol{\varepsilon}$$

$$\mathbf{b} = (\mathbf{b}_1^\top, \dots, \mathbf{b}_n^\top)^\top \quad \& \quad \boldsymbol{\varepsilon} = (\boldsymbol{\varepsilon}_1^\top, \dots, \boldsymbol{\varepsilon}_n^\top)^\top$$

Aim: Prediction of linear aspects $\boldsymbol{\Psi} = \mathbf{K}\boldsymbol{\beta} + \mathbf{Lb}$

Estimable linear aspects

Definition

$$\Psi = \mathbf{K}\beta + \mathbf{L}b \text{ estimable(predictable)} \Leftrightarrow \exists \mathbf{U} : E(\mathbf{U}Y) = E(\Psi)$$

Lemma

$$\Psi = \mathbf{K}\beta + \mathbf{L}b \text{ estimable(predictable)} \Leftrightarrow \exists \mathbf{U} : \mathbf{K} = \mathbf{U}(\mathbf{1}_n \otimes \mathbf{F})$$

Corollary

① $\Psi = \mathbf{L}b \text{ estimable(predictable)} \quad \forall \mathbf{L}$

② $\mathbf{F} \text{ full rank} \Rightarrow \Psi = \mathbf{K}\beta + \mathbf{L}b \text{ estimable(predictable)} \quad \forall \mathbf{K} \text{ \& \ } \forall \mathbf{L}$

Estimable linear aspects

Special cases:

- Population parameter $\Psi = \beta$

estimable if \mathbf{F} full rank $\Leftrightarrow \mathbf{F}^T \mathbf{F}$ regular

- Individual deviations $\Psi = \mathbf{b}_i$

always predictable

- Individual parameters $\Psi = \beta_i = \beta + \mathbf{b}_i$

predictable if \mathbf{F} full rank $\Leftrightarrow \mathbf{F}^T \mathbf{F}$ regular

Prediction of individual effects

- BLUP for individual parameters β_j :

$$\hat{\beta}_i = \mathbf{D}((\mathbf{F}^\top \mathbf{F})^{-1} + \mathbf{D})^{-1} \hat{\beta}_{i;\text{ind}} + (\mathbf{F}^\top \mathbf{F})^{-1}((\mathbf{F}^\top \mathbf{F})^{-1} + \mathbf{D})^{-1} \hat{\beta}$$

where $\hat{\beta} = (\mathbf{F}^\top \mathbf{F})^{-1} \mathbf{F}^\top \bar{\mathbf{Y}}$ & $\hat{\beta}_{i;\text{ind}} = (\mathbf{F}^\top \mathbf{F})^{-1} \mathbf{F}^\top \mathbf{Y}_i$, $\bar{\mathbf{Y}} = \frac{1}{n} \sum_{i=1}^n \mathbf{Y}_i$

- MSE matrix for $\hat{\mathbf{B}} = (\hat{\beta}_1^\top, \dots, \hat{\beta}_n^\top)^\top$:

$$\sigma^2 \left(\left(\mathbf{I}_n - \frac{1}{n} \mathbf{1}_n \mathbf{1}_n^\top \right) \otimes (\mathbf{F}^\top \mathbf{F} + \mathbf{D}^{-1})^{-1} + \frac{1}{n} \left(\mathbf{1}_n \mathbf{1}_n^\top \right) \otimes (\mathbf{F}^\top \mathbf{F})^{-1} \right)$$

Prediction of individual effects

- BLUP for individual deviations \mathbf{b}_i :

$$\hat{\mathbf{b}}_i = \mathbf{D}((\mathbf{F}^\top \mathbf{F})^{-1} + \mathbf{D})^{-1} (\hat{\beta}_{i;\text{ind}} - \hat{\beta})$$

where $\hat{\beta} = (\mathbf{F}^\top \mathbf{F})^{-1} \mathbf{F}^\top \bar{\mathbf{Y}}$ & $\hat{\beta}_{i;\text{ind}} = (\mathbf{F}^\top \mathbf{F})^{-1} \mathbf{F}^\top \mathbf{Y}_i$, $\bar{\mathbf{Y}} = \frac{1}{n} \sum_{i=1}^n \mathbf{Y}_i$

- MSE matrix for $\hat{\mathbf{b}} = (\hat{\mathbf{b}}_1^\top, \dots, \hat{\mathbf{b}}_n^\top)^\top$:

$$\sigma^2 \left(\left(\mathbf{I}_n - \frac{1}{n} \mathbf{1}_n \mathbf{1}_n^\top \right) \otimes (\mathbf{F}^\top \mathbf{F} + \mathbf{D}^{-1})^{-1} + \frac{1}{n} (\mathbf{1}_n \mathbf{1}_n^\top) \otimes \mathbf{D} \right)$$

Optimal designs

Experimental settings x_1, \dots, x_m not necessarily all distinct

- *Distinct settings* x_1, \dots, x_k
- *Numbers of replications* m_1, \dots, m_k ; $\sum_{j=1}^k m_j = m$

Individual design

$$\xi = \begin{pmatrix} x_1, \dots, x_k \\ w_1, \dots, w_k \end{pmatrix}$$

$$w_j = \frac{m_j}{m} \quad \& \quad \sum_{j=1}^k w_j = 1$$

Notation

$$M(\xi) = \frac{1}{m} \mathbf{F}^\top \mathbf{F} \quad \& \quad \Delta = m\mathbf{D}$$

IMSE-criterion

IMSE-criterion for prediction of individual parameters β_j :

$$\Phi_{IMSE,\beta}(\xi) := \int_{\mathcal{X}} \mathbb{E} \left(\sum_{i=1}^n (\hat{\mu}_i(x; \xi) - \mu_i(x))^2 \right) \nu(dx)$$

$$\hat{\mu}_i(x; \xi) = \mathbf{f}(x)^\top \hat{\beta}_i(\xi)$$

$$\Phi_{IMSE,\beta}(\xi) = \text{tr}(\mathbf{M}(\xi)^{-1} \mathbf{V}) + (n-1) \text{tr}((\mathbf{M}(\xi) + \mathbf{\Delta}^{-1})^{-1} \mathbf{V})$$

$$\mathbf{V} = \int_{\mathcal{X}} \mathbf{f}(x) \mathbf{f}(x)^\top \nu(dx)$$

*weighted sum of IMSE-criterion in fixed effects model
 and Bayesian IMSE-criterion*

IMSE-criterion

IMSE-criterion for prediction of individual deviations \mathbf{b}_i :

$$\Phi_{IMSE,b}(\xi) := \int_{\mathcal{X}} \mathbb{E} \left(\sum_{i=1}^n (\hat{\mu}_i^b(x; \xi) - \mu_i^b(x))^2 \right) \nu(dx)$$

$$\mu_i^b(x) = \mathbf{f}(x)^\top \mathbf{b}_i \quad \& \quad \hat{\mu}_i^b(x; \xi) = \mathbf{f}(x)^\top \hat{\mathbf{b}}_i(\xi)$$

$$\Phi_{IMSE,b}(\xi) = \text{tr}(\mathbf{\Delta} \mathbf{V}) + (n-1) \text{tr} \left((\mathbf{M}(\xi) + \mathbf{\Delta}^{-1})^{-1} \mathbf{V} \right)$$

$$\mathbf{V} = \int_{\mathcal{X}} \mathbf{f}(x) \mathbf{f}(x)^\top \nu(dx)$$

Bayesian IMSE-criterion

D -criterion

D -criterion for prediction of individual parameters β_j :

$$\Phi_{D,\beta}(\xi) := \ln \det \left(\text{Cov} \left(\hat{\mathbf{B}}(\xi) - \mathbf{B} \right) \right)$$

$$\mathbf{B} = (\beta_1^\top, \dots, \beta_n^\top)^\top \quad \& \quad \hat{\mathbf{B}}(\xi) = (\hat{\beta}_1(\xi)^\top, \dots, \hat{\beta}_n(\xi)^\top)^\top$$

$$\Phi_{D,\beta}(\xi) = \ln \det (\mathbf{M}(\xi)^{-1}) + (n-1) \ln \det ((\mathbf{M}(\xi) + \mathbf{\Delta}^{-1})^{-1})$$

*weighted sum of D -criterion in fixed effects model
 and Bayesian D -criterion*

D singular: *modified D -criterion*

D -criterion

D -criterion for prediction of individual deviations \mathbf{b} :

$$\Phi_{D,b}(\xi) := \ln \det \left(\text{Cov} \left(\hat{\mathbf{b}}(\xi) - \mathbf{b} \right) \right)$$

$$\mathbf{b} = (\mathbf{b}_1^\top, \dots, \mathbf{b}_n^\top)^\top \quad \& \quad \hat{\mathbf{b}}(\xi) = (\hat{\mathbf{b}}_1(\xi)^\top, \dots, \hat{\mathbf{b}}_n(\xi)^\top)^\top$$

$$\Phi_{D,b}(\xi) = \ln \det (\mathbf{\Delta}) + (n - 1) \ln \det ((\mathbf{M}(\xi) + \mathbf{\Delta}^{-1})^{-1})$$

Bayesian D -criterion

\mathbf{D} singular: *modified D -criterion*

c-criterion

c-criterion for prediction of individual parameters β_i :

$$\Phi_{c,\beta}(\xi) := \sum_{i=1}^n \left(\mathbb{E} \left((\mathbf{c}^\top \hat{\beta}_i(\xi) - \mathbf{c}^\top \beta_i)^2 \right) \right)$$

$$\Phi_{c,\beta}(\xi) = \mathbf{c}^\top \mathbf{M}(\xi)^{-1} \mathbf{c} + (n-1) \mathbf{c}^\top (\mathbf{M}(\xi) + \mathbf{\Delta}^{-1})^{-1} \mathbf{c}$$

*weighted sum of c-criterion in fixed effects model
 and Bayesian c-criterion*

Special case: $\mathbf{c} = \mathbf{f}(x_0)$

interpolation & extrapolation

c-criterion

c-criterion for prediction of individual deviations \mathbf{b}_j :

$$\Phi_{\mathbf{c},b}(\xi) := \sum_{i=1}^n \left(\mathbb{E} \left((\mathbf{c}^\top \hat{\mathbf{b}}_i(\xi) - \mathbf{c}^\top \mathbf{b}_i)^2 \right) \right)$$

$$\Phi_{\mathbf{c},b}(\xi) = \mathbf{c}^\top \mathbf{\Delta} \mathbf{c} + (n-1) \mathbf{c}^\top (\mathbf{M}(\xi) + \mathbf{\Delta}^{-1})^{-1} \mathbf{c}$$

Bayesian c-criterion

Special case: $\mathbf{c} = \mathbf{f}(x_0)$, $x_0 \in \mathcal{X}$ & $x_0 \notin \mathcal{X}$

interpolation & extrapolation

Optimality condition

Approximate design

$$\xi = \begin{pmatrix} x_1, \dots, x_k \\ w_1, \dots, w_k \end{pmatrix}$$

$$w_j \geq 0 \quad \& \quad \sum_{j=1}^k w_j = 1$$

General equivalence theorem

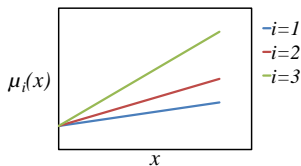


Optimality condition for approximate designs

Example: straight line regression

$$Y_{ij} = \beta_{i1} + \beta_{i2}x_j + \varepsilon_{ij}, \quad x_j \in [0, 1]$$

- $\mathbf{D} = \begin{pmatrix} d_1 & 0 \\ 0 & d_2 \end{pmatrix}$
- small intercept variance d_1



Individual optimal design:

$$\xi_w = \begin{pmatrix} 0 & 1 \\ 1-w & w \end{pmatrix}$$

IMSE- & D -optimal designs

Prediction of

individual parameters β_i

- *IMSE-criterion*

- *D -criterion*

individual deviations \mathbf{b}_i

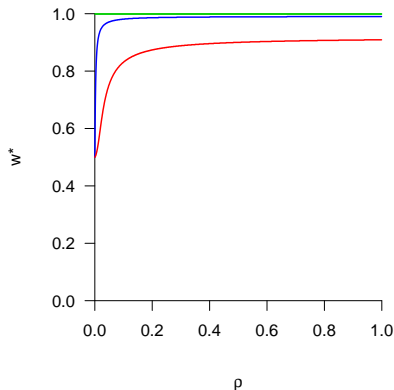
- *IMSE- & D -criterion*

$$\rho = d_2 / (1 + d_2)$$

$$d_1 = 0.001$$

$$n = 100$$

$$m = 10$$



Efficiency of equi-replicated design ($w = 0.5$)

Prediction of

individual parameters β_i

– *IMSE-criterion*

$$\text{eff} = \frac{\Phi_{\text{IMSE},\beta}(\xi_{w^*})}{\Phi_{\text{IMSE},\beta}(\xi_{0.5})}$$

– *D-criterion*

$$\text{eff} = \left(\frac{\exp(\Phi_{D,\beta}(\xi_{w^*}))}{\exp(\Phi_{D,\beta}(\xi_{0.5}))} \right)^{\frac{1}{2n}}$$

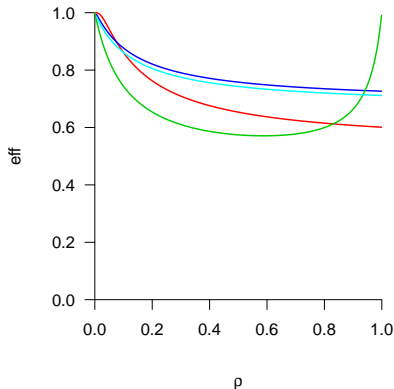
individual deviations b_i

– *IMSE-criterion*

$$\text{eff} = \frac{\Phi_{\text{IMSE},b}(\xi_{w^*})}{\Phi_{\text{IMSE},b}(\xi_{0.5})}$$

– *D-criterion*

$$\text{eff} = \left(\frac{\exp(\Phi_{D,b}(\xi_{w^*}))}{\exp(\Phi_{D,b}(\xi_{0.5}))} \right)^{\frac{1}{2n}}$$



c-optimal designs: interpolation

Prediction of

individual parameters β_i

- $x_0 = 0$
- $x_0 = 0.5$
- $x_0 = 1$

individual deviations b_i

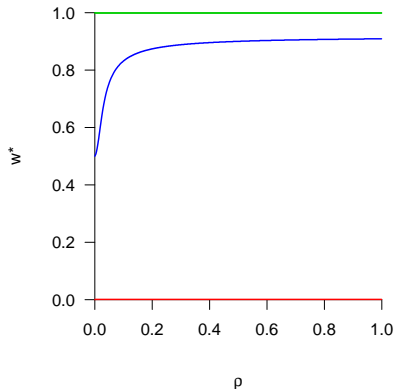
- for all $x_0 \in [0, 1]$

$$\rho = d_2 / (1 + d_2)$$

$$d_1 = 0.001$$

$$n = 100$$

$$m = 10$$



Efficiency of optimal design in fixed effects model ($w = x_0$)

Prediction of

individual parameters β_i

$$eff = \frac{\Phi_{c,\beta}(\xi_{w^*})}{\Phi_{c,\beta}(\xi_{x_0})}$$

- $x_0 = 0.5$

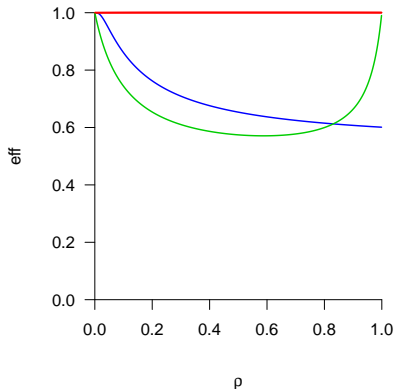
- $x_0 \in \{0, 1\}$

individual deviations b_i

$$eff = \frac{\Phi_{c,b}(\xi_{w^*})}{\Phi_{c,b}(\xi_{x_0})}$$

- $x_0 = 0.5$

- $x_0 \in \{0, 1\}$



c-optimal designs: extrapolation

Prediction of

individual parameters β_i

- $x_0 = 2$

- $x_0 = 10000$

individual deviations b_i

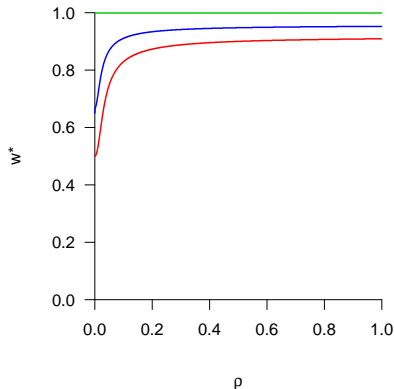
- for all $x_0 > 1$

$$\rho = d_2 / (1 + d_2)$$

$$d_1 = 0.001$$

$$n = 100$$

$$m = 10$$



Efficiency of optimal design in fixed effects model ($\tilde{w} = \frac{x_0}{2x_0-1}$)

Prediction of

individual parameters β_i

$$eff = \frac{\Phi_{c,\beta}(\xi_{w^*})}{\Phi_{c,\beta}(\xi_{\tilde{w}})}$$

- $x_0 = 2$

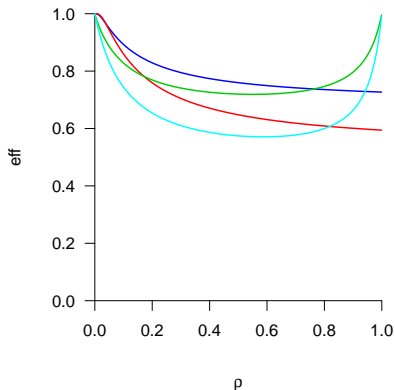
- $x_0 = 10000$

individual deviations b_i

$$eff = \frac{\Phi_{c,b}(\xi_{w^*})}{\Phi_{c,b}(\xi_{\tilde{w}})}$$

- $x_0 = 2$

- $x_0 = 10000$



Outlook

- Different designs for different individuals
- Exact designs
- Unknown dispersion matrix of random effects
- ...

Thank you for your attention!

